

## On the Convexity of a Rate Function

**Choon Peng Tan**

Faculty of Information and Communication Technology, Universiti Tunku Abdul Rahman,  
13 Jalan 13/6, 46200 Petaling Jaya, Selangor, Malaysia  
tanpc@mail.utar.edu.my

**ABSTRACT** Consider a sequence of independent, identically distributed random variables from a common distribution  $Q$ . Suppose that there are  $k$  rare events associated with  $Q$ , for example events involving the  $k$  sample moments. Then the dominant rate at which the probability of the intersection of the  $k$  events converges to zero can be expressed as the relative entropy of a certain distribution and  $Q$ . Considering the intersection of the  $k$  events as a function of  $\alpha$  in a  $k$ -dimensional Euclidean space, we shall show that the dominant rate of convergence is convex in  $\alpha$ . This result is the consequence of the convex property of a certain function associated with the dominant rate of convergence which shall be shown.

**ABSTRAK** Pertimbangkan satu jujukan pembolehubah rawak yang tertabur secara secaman dan tak bersandar daripada satu taburan sepunya  $Q$ . Katalah terdapat  $k$  peristiwa kejarangan yang bersekutu dengan  $Q$ , misalnya peristiwa-peristiwa yang melibatkan  $k$  momen sample. Maka kadar dominant yang mana kebarangkalian persilangan  $k$  peristiwa menumpu kepada sifar boleh diungkapkan sebagai entropi relatif bagi sesuatu taburan dan  $Q$ . Dengan mempertimbangkan persilangan  $k$  peristiwa sebagai satu fungsi  $\alpha$  dalam ruang Euklidan berdimensi  $k$ , kami akan tunjukkan bahawa kadar penumpuan dominant adalah cembung dalam  $\alpha$ . Hasil ini adalah akibat sifat kecembungan sesuatu fungsi lain yang bersekutu dengan kadar penumpuan dominant yang akan ditunjukkan.

(Convex property, dominant rate of convergence, large deviation theory, relative entropy)

### INTRODUCTION

Let  $X_1, X_2, \dots, X_n$  be a sequence of independent, identically distributed random variables from a discrete distribution  $Q(x)$ . Consider  $k$  linearly independent functions of  $X$ , namely  $g_1(X), g_2(X), \dots, g_k(X)$  and define

$k$  events  $\left\{ \frac{1}{n} \sum_{i=1}^n g_1(X_i) \geq \alpha_1 \right\}$ ,

$\left\{ \frac{1}{n} \sum_{i=1}^n g_2(X_i) \geq \alpha_2 \right\}, \dots, \left\{ \frac{1}{n} \sum_{i=1}^n g_k(X_i) \geq \alpha_k \right\}$

associated with these functions, where  $\alpha_1, \alpha_2, \dots, \alpha_k$  do not depend on  $n$ . For

example, if  $g_j(X_i) = X_i^r$  for some  $j$  and integer  $r \geq 1$ , then the  $j$ th event

$$\left\{ \frac{1}{n} \sum_{i=1}^n X_i^r \geq \alpha_j \right\}$$

represents the  $r$ th sample moment larger than the constant  $\alpha_j$ . We are interested in the probability of

$$\left\{ \frac{1}{n} \sum_{i=1}^n X_i^r \geq \alpha_j \right\}$$

where  $\alpha_j > E_Q(X^j)$ . In general, we are interested in the probability of the intersection of the  $k$  rare events, namely,

$$\Pr\left\{\frac{1}{n}\sum_{i=1}^n g_j(X_i) \geq \alpha_j \text{ for } j=1,2,\dots,k\right\}.$$

This probability is equivalent to

$$\Pr\left\{\sum_x P_{X^n}(x)g_j(x) \geq \alpha_j \text{ for } j=1,2,\dots,k\right\},$$

where  $P_{X^n}$  is the type of  $X^n = (X_1, X_2, \dots, X_n)$  and  $P_{X^n}(x)$  is the relative frequency of  $x$  in  $x^n$ , for  $x$  in the alphabet of  $X$ . The evaluation of this probability, denoted by  $Q^n(E)$ , is done by defining a convex set of probability distributions

$$E = \left\{P: \sum_x P(x)g_j(x) \geq \alpha_j, j=1,2,\dots,k\right\} \quad (1)$$

Then by Sanov's Theorem [1] [2],

$$\frac{1}{n} \log Q^n(E) \rightarrow -D(P^* \| Q) \text{ as } n \rightarrow \infty, \quad (2)$$

where  $D(\cdot \| \cdot)$  denotes the relative entropy of two probability distributions,  $P^*$  is the distribution that minimizes  $D(P \| Q)$  over all  $P$  in  $E$ . By using the method of Lagrange multipliers [1], the distribution  $P^*$  is shown to be

$$P^*(x) = cQ(x)2^{\sum_{j=1}^k \lambda_j g_j(x)} \quad (3)$$

where the normalizing constant

$$c = \left[ \sum_x Q(x)2^{\sum_{j=1}^k \lambda_j g_j(x)} \right]^{-1}$$

$\lambda_1, \lambda_2, \dots, \lambda_k$  are positive solutions to

$$\sum_x P^*(x)g_j(x) = \alpha_j, \quad j=1,2,\dots,k, \quad (4)$$

if  $\alpha_j > E_Q[g_j(X)]$ ,  $j=1,2,\dots,k$  and logarithms are to base 2. For large  $n$ ,

$$Q^n(E) \approx 2^{-nD(P^* \| Q)} \quad (6)$$

and hence  $D(P^* \| Q)$  is the dominant rate of convergence of  $Q^n(E)$  to 0. In this paper, we shall assume that  $\alpha = (\alpha_1, \alpha_2, \dots, \alpha_k)$  is a variable and consider  $D(P^* \| Q)$  as a function of  $\alpha$ . We shall show that  $D(P^* \| Q)$  is a convex function of  $\alpha$ . In a previous paper [3], we have studied the properties of  $D(P^* \| Q)$  as a function of the scalar  $\alpha$ .

### MAIN RESULTS

#### Theorem

Let  $g_1(x), g_2(x), \dots, g_k(x)$  be  $k$  linearly independent functions of  $x$  and

$$E = \left\{P: \sum_{x \in A} P(x)g_j(x) \geq \alpha_j, j=1,2,\dots,k\right\}$$

be a convex set of probability distributions, where  $A$  is the alphabet set of the random variable  $X$ . For  $Q \notin E$ , let

$$P_\alpha^*(x) = cQ(x)2^{\sum_{j=1}^k \lambda_j(\alpha)g_j(x)} \text{ for } x \in A$$

be the distribution that minimizes  $D(P \parallel Q)$  over all  $P \in E$ , where  $\lambda = (\lambda_1(\alpha), \lambda_2(\alpha), \dots, \lambda_k(\alpha))$  is the solution to the  $k$  equations

$$\sum_{x \in A} P_{\alpha}^*(x) g_j(x) = \alpha_j, \quad j = 1, 2, \dots, k,$$

the normalizing constant

$$c = \left[ \sum_{x \in A} Q(x) 2^{\sum_{j=1}^k \lambda_j(\alpha) g_j(x)} \right]^{-1}$$

**Proof**

(i)

$$\frac{\partial h}{\partial \lambda_i} = \frac{\left[ \sum_x Q(x) g_i(x) 2^{\sum_{j=1}^k \lambda_j(\alpha) g_j(x)} \right]}{\left[ \sum_x Q(x) 2^{\sum_{j=1}^k \lambda_j(\alpha) g_j(x)} \right]} = \alpha_i, \tag{7}$$

and,

$$\frac{\partial^2 h}{\partial \lambda_i \partial \lambda_j} = (\ln 2) \left\{ E_{P_{\alpha}^*} [g_i(X) g_j(X)] - E_{P_{\alpha}^*} [g_i(X)] E_{P_{\alpha}^*} [g_j(X)] \right\}$$

for  $i = 1, 2, \dots, k; j = 1, 2, \dots, k$ .

(8)

Then,

(i)

$$h(\lambda) = \log \left[ \sum_{x \in A} Q(x) 2^{\sum_{j=1}^k \lambda_j(\alpha) g_j(x)} \right]$$

is convex in  $\lambda$ .

(ii)

$D(P_{\alpha}^* \parallel Q)$  is convex in  $\alpha$ .

The Hessian  $\left( \frac{\partial^2 h}{\partial \lambda_i \partial \lambda_j} \right)$  of  $h(\lambda)$  is  $(\ln 2)\Sigma$ ,

where  $\Sigma$  is the covariance matrix of the random vector  $\mathbf{g}(X) = (g_1(X), g_2(X), \dots, g_k(X))$

and  $X$  has the distribution  $P_\alpha^*$ . Since

$\left( \frac{\partial^2 h}{\partial \lambda_i \partial \lambda_j} \right)$  is positive semidefinite,

therefore  $h(\lambda)$  is convex in  $\lambda$ .

$$\begin{aligned} \text{(ii) } D(P_\alpha^* \parallel Q) &= \sum_x P_\alpha^*(x) \log \frac{P_\alpha^*(x)}{Q(x)} \\ &= \sum_x P_\alpha^*(x) \left[ \log c + \sum_{j=1}^k \lambda_j(\alpha) g_j(x) \right] \\ &= -h(\lambda) + \sum_{j=1}^k \lambda_j(\alpha) \alpha_j \end{aligned} \tag{9}$$

This is equivalent to:

$$h(\lambda_1) - h(\lambda_2) + \sum_{j=1}^k \lambda_j(\alpha_2) \alpha_j(2) - \sum_{j=1}^k \lambda_j(\alpha_1) \alpha_j(1) \geq \sum_{i=1}^k \frac{\partial D(P_\alpha^* \parallel Q)}{\partial \alpha_i} \Big|_{\alpha=\alpha_1} [\alpha_i(2) - \alpha_i(1)]$$

where  $\alpha_i = (\alpha_1(i), \alpha_2(i), \dots, \alpha_k(i))$

and

$$\lambda_i = (\lambda_1(\alpha_i), \lambda_2(\alpha_i), \dots, \lambda_k(\alpha_i))$$

for  $i = 1, 2$ . From (7) and (10), this reduces to

and hence,

$$\frac{\partial D(P_\alpha^* \parallel Q)}{\partial \alpha_i} = - \sum_{j=1}^k \frac{\partial h}{\partial \lambda_j} \frac{\partial \lambda_j}{\partial \alpha_i} + \sum_{j=1}^k \frac{\partial \lambda_j}{\partial \alpha_i} \alpha_j + \lambda_i$$

and from (7),

$$\frac{\partial D(P_\alpha^* \parallel Q)}{\partial \alpha_i} = \lambda_i \quad \text{for } i = 1, 2, \dots, k. \tag{10}$$

From (9), the function  $D(P_\alpha^* \parallel Q)$  is convex in  $\alpha$  if and only if, for any  $\alpha_1$  and  $\alpha_2$ ,

$$D(P_{\alpha_2}^* \parallel Q) - D(P_{\alpha_1}^* \parallel Q) \geq \nabla D(P_\alpha^* \parallel Q) \Big|_{\alpha=\alpha_1} (\alpha_2 - \alpha_1)$$

$$h(\lambda_1) - h(\lambda_2) \geq \sum_{i=1}^k \alpha_i(2) [\lambda_i(\alpha_1) - \lambda_i(\alpha_2)]$$

$$= \sum_{i=1}^k \frac{\partial h}{\partial \lambda_i} \Big|_{\lambda=\lambda_2} [\lambda_i(\alpha_1) - \lambda_i(\alpha_2)]$$

$$= \nabla h \Big|_{\lambda=\lambda_2} (\lambda_1 - \lambda_2).$$

Hence  $D(P_\alpha^* \parallel Q)$  is convex in  $\alpha$  if and only if  $h(\lambda)$  is convex in  $\lambda$ . The result in (i) implies that  $D(P_\alpha^* \parallel Q)$  is convex in  $\alpha$ .

#### REFERENCES

1. Cover, T.M. and Thomas, J.A. (1991). *Elements of Information Theory*. New York: John Wiley.
2. Sanov, I.N. (1957). On the probability of large deviations of random variables. *Mat. Sbornik* 42: 11 - 44.
3. Tan, C.P. (2002). The behaviour of the error probabilities in hypothesis testing for large sample sizes (Malay). *Proceedings of the Tenth National Symposium on the Mathematical Sciences*, Skudai, Malaysia: Penerbit Jabatan Matematik, Universiti Teknologi Malaysia: 436 - 442.